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To move or not to move? Exploring the relationship between residential mobility, risk of cardiovascular disease and ethnicity in New Zealand

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To move or not to move? Exploring the relationship between residential mobility, risk of cardiovascular disease and ethnicity in New Zealand

3 Abstract

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Residential mobility can have negative impacts on health, with some studies finding that residential 4 mobility can contribute to widening health gradients in the population. However, ethnically 5 6 differentiated experiences of residential mobility and the relationship with health are neglected in the 7 literature. To examine the relationship between residential mobility, risk of cardiovascular disease (CVD) and ethnicity, we constructed a cohort of 2,077,470 participants aged 30+ resident in New 8 9 Zealand using encrypted National Health Index (eNHI) numbers linked to individual level routinely recorded data. Using binary logistic regression, we model the risk of CVD for the population stratified 10 by ethnic group according to mover status, baseline deprivation and transitions between deprivation 11 12 statuses. We show that the relationship between residential mobility and CVD varies between ethnic groups and is strongly influenced by the inter-relationship between residential mobility and 13 14 deprivation mobility. Whilst residential mobility is an important determinant of CVD, much of the 15 variation between ethnic groups is explained by contrasting deprivation experiences. To reduce 16 inequalities in CVD within New Zealand, policies must focus on residentially mobile Māori, Pacific and South Asian populations who already have a heightened risk of CVD living in more deprived 17 18 areas.

19 Key words

20 New Zealand; CVD; Ethnicity; Inequalities; Mobility; Migration; Deprivation; Record Linkage

21 Introduction

Cardiovascular disease (CVD) and associated morbidities are among the leading causes of global
deaths (World Health Organisation, 2014). In New Zealand (NZ) there are marked variations between
ethnic groups in the prevalence of CVD (Blakely et al., 2004; Riddell et al., 2007; Jatrana and
Blakely, 2008; Kerr et al., 2008; Grey et al., 2010; Mehta et al., 2011; Perumal et al., 2012; Ker et al.,
2015; Mehta et al., 2014; Exeter et al., 2015; Wells et al., 2015). Between 1980 and 1999, while all

27 ethnic groups experienced reductions in CVD mortality, Māori and Pacific populations saw markedly smaller reductions than non-Māori non-Pacific (nMnP) groups (Blakely et al., 2005). By 2007, these 28 disparities had not disappeared: Maori males and females almost invariably had the highest age-29 specific prevalence of CVD across all age groups, as well as the highest age-standardised prevalence 30 31 of CVD (7.41 compared to NZ's total population at 4.77, and 5.68 for the Pacific group) (Cheuk Chan et al., 2008). Stark differences in risk of CVD and CVD mortality between ethnic groups are not 32 restricted to NZ. For example, rates of ischaemic heart disease amongst South Asian males are 30 to 33 40% higher than rates amongst the UK's general population (Department of Health, 2001). In the US 34 in 2013, Black groups had 30% higher mortality from CVD than Whites, increasing to 113% higher 35 CVD mortality than Asians and Pacific Islanders (Singh et al., 2015). 36

Exploring why ethnic inequalities in CVD exist is therefore of international importance. The existence of these inequalities across different contexts *and* across different ethnic groups suggests that these disparities are not solely explained by 'ethnicity'. Rather, these differences may (in part) be explained by similarities in the experiences of *minority* groups across different contexts and the social gradient to risk of CVD.

The impact of both traditional and environmental risk factors for CVD is modified by socioeconomic status (Albert et al., 2006). Thus, lower socioeconomic status and general disadvantage are associated with higher levels of CVD (Kanjilal et al., 2006; Clark et al., 2009) or increased exposure to CVD risk factors, such as smoking or low levels of physical activity (Gupta et al., 2012). A review of CVD mortality in the US and 11 western European countries found that risk increased with decreasing occupational class and lower levels of educational attainment, as well as factors such as smoking uptake and alcohol consumption (Mackenbach et al., 2000).

Given the social gradient of CVD occurrences, it is important to consider the contrasting socioeconomic circumstances which invariably characterise the experience of marginalised minority ethnic groups (MEGs) in different contexts, particularly when assessing ethnic inequalities in CVD. Where broader structural inequalities exist, these may exaggerate the already disadvantaged

experience of marginalised MEGs and exacerbate health differences. For example, it has been suggested that in NZ, widening inequalities in employment, housing, education and income during the 1980s and 1990s between Māori and Pacific groups compared to non-Māori non-Pacific groups may have had significant health implications (Blakely et al., 2005). This may explain the smaller reductions in CVD mortality for Māori and Pacific populations than observed for the non-Māori non-Pacific population. However, results of a previous study in Auckland, NZ suggest that there is an additional mechanism potentially driving inequalities in CVD: residential mobility.

Exeter et al. (2015) found residential mobility to be an important determinant of CVD in Auckland, 60 NZ. Residential mobility has important implications for health (Morris et al., 2016), and has been 61 examined in NZ in the context of child health outcomes (Jelleyman and Spencer, 2008), but also more 62 63 generally in Australia (Larson et al., 2004) and the UK (Boyle et al., 2005; Norman et al., 2005; 2011). However, the relationship between residential mobility and CVD is under-explored. In 64 particular, no previous work has specifically investigated whether this relationship varies by ethnic 65 group. Residential mobility is an inherently selective event: a wealth of research demonstrates this, 66 67 highlighting that movers are often distinct from stayers in their age, sex, stage in the lifecourse, tenure, educational attainment, social class, income and health (e.g. Bentham, 1988; Findlay, 1988; 68 Simpson and Finney, 2009). As the socioeconomic circumstances of different ethnic groups in any 69 70 socio-political context varies, with substantial evidence that people from ethnic minorities also have 71 significantly worse health experiences than people from non-ethnic minority groups, the patterning to residential mobility may vary between ethnic groups. More importantly, the nature of residential 72 73 mobility experienced by different ethnic groups may also vary and therefore differently influence risk 74 of CVD. For example, if certain groups are more likely to move frequently over shorter distances, or 75 perhaps move frequently within similarly deprived neighbourhoods, the influence of these moves on CVD risk may vary compared with groups who move infrequently or experience upwards deprivation 76 77 mobility, moving from more to less deprived areas. Results of Exeter et al.'s (2015) research support this, revealing that those moving from less to more deprived areas having a higher risk of CVD 78 79 hospitalisation than those moving in the opposite direction. The concept of health-selective migration

can help us begin to disentangle possible variations in the patterning to residential mobility fordifferent ethnic groups.

82 Theories of health-selective migration hypothesise that health gradients are widened as differently healthy groups of people are sorted into different area types (e.g. Boyle, 2004; Norman et al., 2011; 83 Exeter et al., 2011). Those in good health or with favourable health-related individual characteristics 84 are more likely to experience upward mobility, moving to less deprived areas. Conversely, those in 85 86 poor health or with unfavourable health-related individual characteristics are more likely to experience downward mobility or remain in more deprived areas. These scenarios exacerbate existing 87 health gradients as those in poor health continue to suffer the deleterious consequences of their 88 relative disadvantage, while those living in more advantaged circumstances continue to reap the 89 90 health benefits of their elevated situation. In a recent review of the literature on health and mobility, Morris et al. (2016) distinguish between population level aggregate studies, those which are typically 91 used in the context of discussions of health-selective migration and changing health gradients (e.g. 92 Boyle and Norman, 2009), and individual level studies wherein the relationship between health and 93 94 mobility is more often viewed negatively (e.g. Jelleyman and Spencer, 2008).

95 Thus, in this study we might hypothesise that through health-selective migration, risk of CVD is lower for movers as compared to stayers as those at risk of CVD are less likely to move. However, we might 96 97 also assume that risk of CVD is higher for an individual who has moved due to the stress associated with a move, perhaps exacerbated or attenuated by the nature of the move itself. Moreover, are they 98 moving to a more or less deprived area? Given the results of the previous study (Exeter et al., 2015), 99 we can hypothesise that movers across NZ will also have a higher risk of CVD than stayers, as found 100 in Auckland. However, what is of interest is why this occurs, and whether the relationship varies 101 102 between ethnic groups. This focuses attention on the complex relationship between mobility and health, and the context within which different ethnic groups live out their day-to-day lives. 103

The persistent (albeit narrowing) inequalities in areas such as housing and education experienced by
 MEGs in NZ (see Blakely et al., 2005) are echoed in the overwhelming concentration of minority

106 groups in the most deprived areas of the country (see Table 2). The marginalisation of these groups 107 both spatially but also more broadly (see work on the relationship between poor health outcomes and racial discrimination in NZ such as Harris et al., 2006; Harris et al., 2012; Harris et al., 2015) suggests 108 109 that MEGs in NZ might be more likely to experience increased rates of residential mobility. The 110 neglected concept of 'malign migration' holds that marginalised, socially disadvantaged groups are more likely to experience residential mobility, and this is more common in inner city (often deprived) 111 areas: this is detrimental to health (Warfa et al., 2006). It therefore seems likely that different ethnic 112 groups in NZ will have different experiences of residential mobility, perhaps through processes of 113 'malign migration' but also more broadly in terms of socioeconomic inequalities and the selective 114 nature of migration. We can assume that this will differently influence the relationship between CVD 115 and residential mobility for different ethnic groups. One aspect of the relationship between residential 116 117 mobility and health which gets less specific coverage in the literature is *immobility*. Notwithstanding a few notable exceptions (e.g. Boyle et al., 2004; Exeter et al., 2011; Brown et al., 2012), much of the 118 extant literature in this area focuses on the selection of *mobile* groups into different socioeconomic 119 circumstances. However, reasons for immobility may be as important in the selection process as 120 reasons for mobility. This will also be addressed. 121

This paper uses a unique, unrivalled longitudinal dataset to investigate an under-explored determinant of CVD, that of residential mobility, and evaluate whether the salience of residential mobility (and immobility) as a determinant of CVD varies between ethnic groups. Extending the research for the Auckland Region by Exeter et al. (2015), a cohort of participants are derived from national routine health databases in NZ. We address the following research questions:

- 127 1. Do movers in NZ have a higher risk of CVD than stayers?
- 128 2. Is risk of CVD for movers attenuated by baseline deprivation at the start of the study period?
- 3. Do the patterns observed for movers and stayers in NZ overall vary for specific ethnicgroups?
- 4. How does the nature of a move influence risk of CVD for different ethnic groups in NZ? and;
- 132 5. Does risk of CVD for ethnic groups who do not move (stayers) vary by deprivation?

133 Data and methods

A cohort of participants was identified using the unique health identifier which is assigned to the majority of all NZ residents. Using these identifiers, patient records are anonymously and securely linked between four national routine health databases: enrolment with a Primary Health Organisation (PHO), hospital discharges, mortality records and pharmaceutical dispensing claims from community pharmacies. As data held by the Ministry of Health on discharges from private hospitals are incomplete, these are excluded from the cohort (Ministry of Health, 2014).

Building on Exeter et al.'s (2015) study, we use the same population eligibility criteria, but increase the coverage to the entire adult population of NZ rather than focusing on Auckland residents. Thus, participants are eligible for inclusion if enrolled in any PHO within NZ during at least one of the 34 calendar quarters of the study period from 1 January 2006 to 30 June 2014; aged 30 years or over at the start of the study period; had complete demographic information; and had no prior history of CVD (defined below) before 1 January 2006. Figure 1 summarises the eligibility criteria for this study.

146

FIGURE 1 ABOUT HERE

147 Variables

Variables identifying each participant's age, sex, ethnicity and area of residence are the key independent demographic variables for this analysis. Consistent with previous work, age was categorised into six groups (30-44; 45-54; 55-64; 65-74; 75-85) with the 55-64 age band used as the reference group (Exeter et al., 2015; Grey et al., 2014; Warin et al., 2016). The age group was restricted due to the low risk of CVD for those aged below 30 years, and the incomplete data, increased risk of having a history of CVD and the statistical problem of small numbers for those aged over 85.

Using the national ethnicity coding protocols for NZ, we prioritised ethnicity to identify five ethnic groups: Māori, Pacific, Indian (Indian groups are distinguishable from Other South Asian groups in NZ's ethnicity coding system), Other Asian, and NZ and Other European combined (NZEO).

Consistent with the PREDICT study (Wells et al., 2015), we distinguish between Indian and Other Asian groups given the higher risk of CVD amongst Indian participants relative to Other Asian participants (Ministry of Health, 2012). We use Census Meshblocks (MBs) to identify a participant's area of residence in each calendar quarter, and to derive information on residential mobility and area deprivation.

MBs consist of (on average) approximately 100 persons and are the most detailed geographic unit of analysis available for census data in NZ. Using the NZ Index of Deprivation (NZDep2006), we assigned a deprivation score to each participant based on their MB for each calendar quarter. This is a measure of area level socioeconomic deprivation based on nine variables from the 2006 Census (Salmond et al., 2007). Scores are ranked into quintiles where quintile 1 (Q1) comprises the least deprived 20% of areas across NZ and quintile 5 (Q5) the most deprived 20%.

169 By assigning each participant to a MB and NZDep2006 score at each calendar quarter, we identified 170 participants who moved during the study period as well as their deprivation trajectory according to moves between or within deprivation quintiles. We focus on overall deprivation trajectory; for 171 172 participants who moved, we investigate the change between first and last recorded MB and NZDep2006 score. We use the same measure of deprivation for all time points (from 2006 to 2014), 173 as NZDep2013 was not published when we obtained our dataset. However, we do recognise that 174 175 areas can change their level of deprivation over time (Norman, 2010), and that changing and persistent area deprivation can have a concomitant influence on health (Boyle et al., 2004; Norman et 176 al., 2010; Exeter et al., 2011). The implications of using fixed deprivation levels to analyse changes in 177 health has been considered elsewhere and found not to affect interpretations (Bajekal et al., 2013). In 178 the main this is because the relative position of areas with regard to their level of deprivation has great 179 consistency over time (Norman and Darlington-Pollock, 2016). 180

Any participant with a previous hospitalisation or procedure related to acute coronary syndrome, ischaemic and haemorrhagic stroke, peripheral arterial disease or for congestive failure was defined as having a CVD event, either for exclusion purposes or for identification during the study period. Table

184 1 summarises the variables included in the analysis, distinguishing between movers and stayers for the
185 NZ cohort of participants.

186 Analysis

We used binary logistic regression to model risk of CVD for different ethnic groups in NZ. All results 187 are expressed as odds ratios (ORs) and accompanied by 95% confidence intervals (CIs). We 188 constructed five models adjusting for: 1) mover status; 2) mover status and baseline deprivation; 3) 189 deprivation mobility status; 4) detailed deprivation transitions; and 5) deprivation circumstances for 190 stayers. Deprivation mobility status identifies the overall nature of the deprivation mobility 191 192 experienced by each participant- moving to more deprivation; churning within comparable deprivation; or moving to less deprivation. The detailed deprivation transitions expand on this, in 193 particular identifying moves into, out of or within the least (Q1) and most (Q5) deprived areas, as well 194 195 as those who move within Q2 to Q4. Given the anticipated role of deprivation in contributing to risk 196 of CVD, the results begin with a discussion of the ethnic profile of the deprivation quintiles (according to baseline deprivation). In the first instance, all models were run using the total sample 197 population, adjusting for age, sex and ethnicity. Then, the five models were stratified by ethnic group, 198 199 adjusting for age and sex (models 1e to 5e). For the models adjusting for stable deprivation, movers 200 are the reference group. For all other models, we use stayers as the reference group in the relevant 201 variables. We take females and NZEO as the reference group for gender and ethnicity. As mentioned above, we take those aged 55-64 as the reference group in line with wider literature investigating 202 203 CVD (e.g. Warin et al., 2016). The models were stratified by ethnic group as we hypothesised that the relationships between residential mobility and risk of CVD may vary by ethnic group. Ethnic-specific 204 205 models illuminate how the relationship between residential mobility and risk of CVD may interact differently with different ethnic groups: this is not captured in models only adjusting for ethnicity. 206 Results for the ethnic-specific models are presented as modelled probabilities. Modelled probabilities 207 208 are more comparable than ORs which only summarise the constant effect of the predictor variable 209 (e.g. becoming less deprived) on risk of CVD. Modelled probabilities quantify the likelihood of CVD

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- 210 for the predictor variable (e.g. becoming less deprived), holding all other variables constant. All
- analyses were conducted in IBM SPSS Statistics 23.

TABLE 1 ABOUT HERE

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214 Results

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i) Ethnic profile of deprivation quintiles in NZ

Table 2 summarises the distribution of each ethnic group across the baseline deprivation quintiles. 216 Māori and Pacific peoples, and to a lesser extent Indians, are disproportionately represented in the 217 more deprived quintiles (Q4 and Q5). For Māori and Pacific, this accounts for the majority of the 218 population. NZEO peoples are skewed towards the less deprived quintiles (Q1-Q3) whilst Other 219 Asian peoples are fairly evenly distributed between Q1 and Q4. Given the unequivocal relationship 220 between poor health and increasing deprivation (e.g. Boyle et al., 2005), the distribution of NZ's 221 222 population across the deprivation quintiles will be pertinent to experiences of specific health 223 outcomes, including CVD.

224

TABLE 2 ABOUT HERE

225 *ii)* The influence of mobility on CVD in a national health database cohort

We summarise the results of each model first for all persons, and then by ethnic group. Table 3 226 presents ORs and CIs for the five all-person models. Statistically significant ORs are starred. Males 227 consistently have significantly higher odds of CVD than females. Adjusting for different residential 228 mobility or deprivation mobility variables has only a marginal impact on the size of the ORs for 229 males. A clear age-gradient in CVD risk is apparent across all models, whereby participants aged 30-230 231 44 and 45-54 years have significantly lower odds of CVD than participants aged 55-64. This reverses 232 in the older age groups: those aged 65-74 and 75-85 years have a significantly higher risk of CVD than the reference group. As with the ORs for gender, adjusting for different residential mobility or 233 deprivation mobility variables has only a marginal impact on the ORs for each age group. This does 234 235 not affect the statistical significance of the variables, or the interpretation of the ORs.

236

TABLE 3 ABOUT HERE

238 Adjusting for residential or deprivation mobility has a more discernible impact on the ORs for certain 239 ethnic groups. Across all five models, the highest odds of CVD are consistently observed for Māori groups, ranging from an OR of 2.26 (95% CI 2.21-2.30) in model 3 to 1.97 (1.93-2.01) in model 2. 240 The odds of Māori having CVD, however, are attenuated by baseline deprivation, evident in the 241 242 reduction of the odds of CVD for Maori in model 2 compared to the other models. Models 1, 3, 4 and 5 all suggest that the odds of Māori being hospitalised for CVD is more than twice that of NZEO. 243 However, when adjusting for baseline deprivation the odds are significantly lower (1.97). The 244 importance of baseline deprivation in explaining odds of CVD is not limited to Māori, as the odds of 245 CVD also notably declines for Pacific and Indian participants in model 2. Baseline deprivation 246 appears to exert a stronger influence on odds of CVD for each ethnic group than mover status alone. 247 Indeed the ORs for each deprivation quintile are all significantly different from each other, increasing 248 249 in size with increasing deprivation with O2 at 1.14 (1.12-1.16) and O5 climbing to 1.57 (1.54-1.59). Odds of CVD for Māori and Pacific groups are more notably attenuated when adjusting for 250 deprivation than the other ethnic groups. It is possible this is largely driven by the likelihood of Māori, 251 Pacific, and to a lesser extent, Indian groups, living in more deprived areas as CVD is socially graded. 252

253 Results of models 4 and 5 further demonstrate the importance of deprivation in explaining risk of CVD for different ethnic groups. ORs are attenuated when adjusting for detailed deprivation 254 255 transitions (model 4) and stable deprivation for stayers (model 5). Although the reduction in the ORs 256 for each ethnic group is smaller in models 4 and 5 than observed in model 2, it is still notable. Despite 257 the apparent importance of deprivation, it is important to note that even after adjusting for deprivation and deprivation transition, the odds of CVD for Māori and Pacific groups are still notably high. 258 Variables not adjusted for in these models, such as social class, tenure, education and employment 259 260 may explain some of the variation observed here. The importance of these variables in relation to risk 261 factors for CVD has been determined in the wider literature (e.g. Albert et al., 2006).

After Māori, Pacific people have the highest odds of CVD, followed by Indians. These three ethnic groups consistently have significantly higher odds of CVD than NZEO, whether adjusting for residential or deprivation mobility. Conversely, Other Asian peoples have significantly lower odds of

265 CVD relative to NZEO in all five models. While the ORs for Māori, Pacific and Indian peoples are 266 attenuated when adjusting for residential or deprivation mobility, this is not true for Other Asians. The 267 odds of Other Asians being hospitalised for CVD are consistently about 45% less likely than for 268 NZEO participants.

In models 1 and 2, movers have significantly higher odds of CVD than stayers (1.26 (1.24-1.27) when 269 adjusting for baseline deprivation). There is no change in the size of the ORs or the size of the 270 271 confidence interval between these two models. The influence of residential mobility on the odds of being hospitalised for CVD can also be seen in model 3: after adjusting for deprivation mobility 272 status, the odds of CVD are significantly higher for movers regardless of their deprivation mobility 273 status. Further, the odds of CVD for these differently mobile groups are not significantly different 274 275 from each other. However, as demonstrated in model 4, the odds of CVD are influenced by detailed deprivation transition: variations begin to emerge when looking at residential mobility in the context 276 of transitions into and out of the extremes of the deprivation spectrum. 277

Movers who churn within the least deprived quintile (Q1) are the only mobile group to have 278 significantly lower odds of CVD than stayers (0.88 (0.85-0.91)). Model 4 shows that the odds of CVD 279 generally increases successively with each transition down the deprivation spectrum. Of those moving 280 within the same deprivation quintile (i.e. churning), the highest odds of CVD are for those churning 281 282 within the most deprived quintile (Q5) (1.71 (1.66-1.76)), followed by those who move out of or into O5. There is no significant difference in the odds of CVD among those moving into O5 (1.52 (1.48-283 1.56)) or out of Q5 (1.55 (1.51-1.58)), or between those moving into (1.08 (1.05-1.11)) or out of (1.06 284 (1.03-1.08)) Q1. 285

Model 5 further demonstrates that movers are, generally, at significantly higher risk of CVD than stayers. Odds of CVD for stayers (in model 5) are consistently significantly lower than for the reference group of movers. Here, we see a clear deprivation gradient with the odds of CVD increasing significantly for stayers with increasing levels of area deprivation. However, despite these significant

increases stayers in Q5, the most deprived area, are still significantly *less* likely than movers to haveCVD.

292 The results of the all-person models suggest: a) there is an important relationship between residential mobility and CVD but that the overall direction of the move is less important than the move itself, and 293 294 b) CVD is socially graded. This is apparent in the clear deprivation gradient in odds of CVD by baseline deprivation, stable deprivation (for stayers), and when accounting for specific moves into and 295 296 out of the most and least deprived areas. Importantly, we also see clear and consistent disparities in the odds of CVD by ethnic group, each somewhat attenuated by residential mobility and deprivation 297 (change). The following set of results explore the social gradient to CVD and the influence of 298 residential mobility and deprivation (change) in more detail for each ethnic group. 299

300 *iii) Ethnic-specific influences of mobility on CVD*

For models 1*e* to 5*e* (subset by ethnic group), modelled probabilities of CVD are calculated for each ethnic group by origin deprivation, deprivation mobility status, detailed deprivation transitions, and stable deprivation for stayers. These are compared to the modelled probabilities of CVD for the total population. All probabilities are derived from models adjusting for age and sex in addition to the relevant residential mobility or deprivation-related variables. Probabilities derived from the allpersons models discussed above also adjust for ethnicity. Error bars are presented on each graph to represent the 95% confidence intervals.

Figure 2 presents the modelled probability of CVD by mover status stratified by ethnicity from models 1*e*. For all ethnic groups, the probability of CVD is significantly higher for movers than for stayers. Compared to the total population, Māori and Pacific movers and stayers, and Indian movers have significantly higher probabilities of CVD. Probability of CVD for Other Asian stayers is significantly lower than the probability of CVD for all other groups (3.31% compared to 17.47% for Māori movers).

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FIGURE 2 ABOUT HERE

316 Figure 3 summarises results from models 2e: the probability of having CVD by baseline deprivation stratified by ethnic group. Whilst a deprivation gradient is apparent for all ethnic groups, the steepness 317 of this gradient varies. It is steepest for Maori and Pacific groups who have a disproportionate share of 318 their population in the more deprived quintiles (see Table 2). Further, although increasing deprivation 319 is generally associated with increasing probabilities of CVD for all groups, Māori groups in Q1-Q5 320 (9.76% - 16.38%), Pacific groups in O1-O5 (7.91% - 10.81%) and Indian groups in O1 (6.24%) have 321 a higher probability of CVD than observed for corresponding quintiles of the NZEO population. 322 323 Differences are significant for Māori. The distribution of probability of CVD by deprivation is flatter around Q2-Q4 for Other Asian, Indian and Pacific groups than for the total population, or for Māori 324

and NZEO groups.

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315

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FIGURE 3 ABOUT HERE

The patterning to probability of CVD varies somewhat between ethnic groups according to their deprivation mobility status (figure 4). For Māori and Pacific groups (18.42% and 14.01% respectively), the highest probability of CVD is for movers who churn within the same deprivation quintile. Differences are significant for Māori. Conversely, for all other ethnic groups movers churning within the same deprivation quintile tend to have lower probabilities of CVD than those who either become more or less deprived, significantly lower for NZEO.

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FIGURE 4 ABOUT HERE

This likely reflects the high concentrations of Māori (68.0%) and Pacific (67.5%) populations residing in Q4 and Q5 at baseline: the majority of their moves will therefore be within very deprived areas. Differences in the probability of CVD between those whose areas become more or less deprived are small for all ethnic groups (less than 0.5% for all groups).

To further explore how the nature of a move influences probability of CVD between ethnic groups,
we also adjusted for detailed deprivation transitions (models 4*e*). Māori groups consistently have the

341 highest probability of CVD when compared to all other ethnic groups in comparable circumstances. 342 There is a significant marked gap between those churning within Q5 (the most deprived quintile) and all other movers within NZEO, Indian and Māori groups (figure 5). Conversely, differences between 343 Other Asian and Pacific groups are much smaller (although still significant for Pacific groups). Indian 344 345 and Other Asian stayers had the lowest probability of CVD compared to mobile Indian or Other Asian peoples. Māori stayers have a higher probability of CVD (14.50%) than Māori movers moving across 346 (significant difference for this group), into or out of the least deprived quintile (9.56%, 13.41% and 347 13.81%, respectively). However, this is unsurprising given that 68.7% of Māori stayers remain in Q4 348 and O5. Pacific and NZEO stayers also have a higher probability of CVD than those moving across, 349 into or out of Q1, but differences are small (but significant for NZEO). It is important to note that as 350 only 4.5% of Pacific reside in Q1 (at baseline) compared to 26.5% of NZEO, the reasons for these 351 352 similar probabilities will vary.

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FIGURE 5 ABOUT HERE

Figure 6 illustrates the results of models 5e as probabilities of CVD by experience of deprivation for 354 stayers compared to movers, stratified by ethnic group. The similarities in the patterning of health for 355 356 stayers by deprivation quintile and for movers by baseline deprivation quintile are striking. The steepest gradient is observed for Māori stayers (differences between quintiles are generally 357 significant). Probability of CVD for Māori stayers who remain in Q5 (16.31%) is more than 1.5 times 358 that of Māori stayers who remain in Q1 (9.17%). However, probability of CVD for stayers in Q5 is 359 360 not significantly different from movers. Conversely, the gradient for Pacific, Indian and Other Asian stayers is less marked with probability of CVD only about 1.2 times greater for stayers in Q5 than for 361 stayers in O1. Movers for these groups consistently have a significantly higher probability of CVD 362 than stayers, irrespective of deprivation. The lowest probabilities of CVD for stayers are consistently 363 found for those remaining in the least deprived areas for all ethnic groups. 364

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FIGURE 6 ABOUT HERE

367 Discussion

This paper aimed to investigate the relationship between residential mobility and risk of CVD for different ethnic groups, building on previous results of a study of Auckland's adults. We expanded the research, exploring whether the relationship between residential mobility and CVD varies between ethnic groups across the whole of NZ. Further, we addressed the role of immobility in explaining differences in health between ethnic groups, an idea that has not been extensively explored in comparable literature.

The key findings of this paper are a) movers have a higher risk of CVD than stayers across the adult 374 population of NZ (similar to the results of Exeter et al.'s (2015) for Auckland's adults); the influence 375 376 of residential mobility on risk of CVD gains in importance through its relationship with deprivation mobility; and c) the relationship between residential mobility and risk of CVD varies notably between 377 ethnic groups. Interpretation of the all-person models (see Table 3) suggested that the salience of 378 residential mobility varied for each ethnic group through the complex relationship with deprivation, 379 380 whether at baseline or through changing deprivation trajectories. Adjusting for baseline deprivation, deprivation mobility status or detailed deprivation transitions attenuated the odds of CVD for all 381 ethnic groups, apart from Other Asians. The importance of deprivation was also apparent in the clear 382 383 gradient to odds of CVD for stayers by deprivation quintile (model 5).

To explore the attenuation of the odds of CVD by ethnic group observed in models 1-5, we calculated 384 385 modelled probabilities of CVD, sub-setting each of the models by ethnic group. We refer to the results 386 of these models as 1e to 5e. Calculating modelled probabilities allows comparisons within and between ethnic groups and reveal a more nuanced picture of the relationships between residential 387 mobility, deprivation and CVD for different ethnic groups in NZ. As with the all-person models, we 388 389 found that movers consistently have a significantly higher probability of CVD than stayers for all ethnic groups. This is consistent with wider literatures investigating the relationship between 390 residential mobility and health (albeit not ethnically differentiated): at the individual level, Morris et 391 392 al. (2016) note that residential mobility is often associated with poorer health outcomes for movers

compared to stayers (see Jelleyman and Spencer, 2008; Scanlon and Devine, 2001; Piro et al., 2007).
However, the nature of the residential mobility event will vary markedly between ethnic groups:
disadvantaged groups will have very different motivations and opportunities for residential mobility
to those of advantaged groups. This, in turn, will influence the relationship with CVD.

To effectively disentangle these relationships, we should look to the detailed health, social and physical histories of individuals. Morris et al. (2016: 2) advocate such an analytical framework, also drawing on individual experiences and personal biographies. Within the scope of this study, we use baseline deprivation and deprivation change (measured as deprivation mobility status and detailed deprivation transitions) to try and unpack the relationship between residential mobility and CVD for different ethnic groups.

403 In the Auckland study, the odds of CVD were lower for those moving up the deprivation spectrum (to lower deprivation) compared to those moving down (to more deprivation). Exeter et al. (2015) 404 question whether health status is more associated with an individual's current residence, or where they 405 406 have been. However, it is more complex than that. We must also examine whether the extent of the influence of current or previous residence varies by, for example, deprivation, and consider the 407 relationship with literatures on selective sorting (see Norman et al., 2005). In terms of the results in 408 Auckland, we might assume that movers take some of the health advantage of more prosperous areas 409 410 with them when moving from less to more deprived areas, while those moving out of more deprived areas may inherit the health status of the less deprived areas they move to, particularly if those groups 411 of movers have been sorted into less deprived areas by virtue of their better health. 412

413 Our results reveal a more nuanced picture for different ethnic groups across NZ, and one with 414 marginal differences when looking at the population as a whole. Maori, Pacific and NZEO movers 415 who move to less deprived areas have a (marginally) higher risk of CVD than their peers moving to 416 move deprived areas, perhaps suggesting they inherit the health status of the areas they move to *or* are 417 sorted into these less deprived areas due to their good health. However, differences between the 418 mobile groups are too small to be significant. Conversely, Indian and Other Asian movers who

419 become more deprived have a higher probability of CVD than their peers who become less deprived. 420 Are these down the deprivation spectrum precipitated by poor health? This downward deprivation 421 mobility is the most detrimental to Indian and Other Asian groups as this is associated with the highest probability of CVD. Yet for Maori and Pacific movers, the highest probability of CVD is 422 423 associated with churning within the same deprivation quintile. Indeed for Maori, churning with the same level deprivation results in significantly higher probabilities of CVD than for any other 424 deprivation mobility status. In contrast, churning within the same level of deprivation for NZEO 425 movers results in a significantly lower probability of CVD. This likely reflects the markedly higher 426 concentration of Maori and Pacific groups in the most deprived quintiles (see Table 2; Salmond and 427 Crampton, 2012): the health of those churning within these deprived areas will likely be poorer than 428 those who have spent time in less deprived areas and then moved down. 429

These results highlight the importance of looking, insofar as possible, to the wider experiences of 430 431 differently mobile groups in order to understand the relationship with risk of CVD. Results of models 4e further illustrate this: Maori and Pacific movers who move within, into or out of the least deprived 432 quintile (O1) all have a lower probability of CVD than their stable counterparts, significantly lower 433 for those moving within Q1. Similarly, NZEO movers churning with Q1 also have a significantly 434 435 lower probability of CVD than their stable counterparts. This strengthens the conclusions drawn above: the health advantage of those groups in Q1 likely reflects their relatively social advantage, here 436 defined by residency in the least deprived quintiles. Maori and Pacific groups residing in the least 437 deprived quintiles will be particularly advantaged compared to their stable peers given the 438 overwhelming concentration of these ethnic groups in the most advantaged areas. 439

It seems likely that deprivation histories interact with the opportunities for residential mobility and the nature of the move itself (in terms of changing deprivation). We must therefore ask, are there different causal pathways operating which might be explaining these results and the marked (often significant) variations within and between ethnic groups?

444 Firstly, those MEGS which concentrate in more deprived areas may have a heightened risk of CVD, irrespective of any residential mobility or the nature of the move, as CVD is socially graded. Those 445 living in socially deprived areas may also be individually deprived, perhaps with lower levels of 446 educational attainment and working in lower occupational classes. Each are associated with a higher 447 448 risk of CVD mortality (Mackenbach et al., 2000): lower educational attainment may mean individuals are less able to participate in health promotion activities or are less aware of appropriate life-style 449 choices and health-enabling behaviours (Glymour et al., 2014). However, those living in more 450 deprived areas may also have access to fewer facilities or services which promote health-enabling 451 behaviours, thus contributing to an increased risk of CVD. These compositional and contextual factors 452 may collectively contribute to ethnic and social disparities in CVD. 453

Secondly, residential mobility is associated with poorer health outcomes as already noted, and this is 454 consistent across ethnic groups. However, the relationship varies, evidenced by the ratio of the 455 456 probabilities of CVD for movers compared to stayers in models 1e: probability of CVD is 1.5 times as likely for NZEO movers compared to stayers, this increases to 1.8 times as likely for Other Asians 2.6 457 times as likely for Indians, and more than 3 times as likely for Maori and Pacific movers. This may be 458 explained by their contrasting deprivation experiences and the extent to which this determines the 459 460 nature of the move itself. To understand this, we must revisit the concept of 'malign migration' and the notion that marginalised, socially excluded groups in inner city, deprived areas "experience higher 461 than average levels of residential mobility which is detrimental to health" (Warfa et al., 2006: 504). 462 26% of the Maori population who moved during the study period moved more than 4 times within the 463 most deprived areas. This increases to 37% of Pacific movers, yet only accounts for 4% of NZEO 464 movers. The interaction between deprivation and higher than average levels of residential mobility 465 may be particularly pertinent to our understanding of the causal pathways driving the varying 466 relationships between residential mobility and CVD for ethnic groups through uptake of health-related 467 behaviours and the relationship with access to healthcare. 468

469 Increased residential mobility is associated with increased participation in risk behaviours, including470 smoking, alcohol consumption even drug use (see Morris et al., 2016 for a review of relevant

471 literatures): these risk factors, particularly smoking, may influence risk of CVD. Participation in these
472 health-related behaviours is socially graded and varies between ethnic groups: while relative
473 deprivation is the most important predictor of smoking uptake in NZ, increased inequality between
474 Maori and non-Maori groups leads to higher smoking rates amongst Maori (Barnett et al., 2005).

475 Residential mobility, particularly amongst those concentrated in more deprived areas, may disrupt 476 access to preventative healthcare services (see Warfa et al., 2006; Jelleyman and Spencer, 2008). However, it is likely that there are additional salient interactions. Healthcare provision has famously 477 478 been found to follow an inverse care law (Hart, 1971) whereby services are inversely distributed 479 according to need. In NZ, recent research concluded that despite improvements in cardiac interventions, the inverse care law in the context of ischaemic heart disease persist for the Maori 480 population (Sandiford et al., 2015: 974). Ethnic differences in access or utilisation of healthcare may 481 be variously explained by cultural, linguistic or religious factors influencing perceptions of healthcare 482 483 services (e.g. willingness or perceived ability to access services) and participation in health promotion activities (Zanchetta and Poureslami, 2006). However, these barriers extend past patient-484 level characteristics, including factors such as the attitudes of healthcare providers or structural 485 barriers in the organisation of the healthcare system (see Scheppers et al., 2006). 486

We might therefore assume that the higher risk of CVD for MEGs churning with more deprived areas can, in part, be explained by the interaction between deprivation, residential mobility (or perhaps 'malign migration'), ethnicity and access to preventative healthcare. Each are associated with a heightened risk of CVD, and collectively reflect a significant policy concern. To extent Jelleyman and Spencer's (2008) arguments in the context of child health outcomes, CVD preventative healthcare services should be reoriented to effectively engage residentially mobile Maori, Pacific and Indian populations living in more deprived areas already vulnerable to CVD.

494 Notwithstanding the likely important of the interactions outlined above, the reported results may be 495 confounded by cultural factors differently influencing the patterning of residential mobility between 496 ethnic groups, or by ethnically differentiated experiences of tenure and housing conditions across NZ.

497 Firstly, despite broad similarities important differences in the age profile of movers across ethnic groups have been observed in the UK (Finney and Simpson, 2008; Simpson and Finney, 2009). 498 Although younger adults are consistently the most mobile, South Asian groups are less likely to move 499 500 than other ethnic groups. Finney and Simpson (2008) attribute this to differences in household 501 formation as South Asian young adults are more likely to remain the family home until marriage contrasting with non-South Asian young adults who are more likely to live alone before marriage. It is 502 reasonable to assume that patterns of residential mobility may be similarly influenced by different 503 cultural traditions in the NZ population which may be pertinent. 504

505 Secondly, recent research has shown that falls in owner-occupied housing have been greater in Maori (20%) and Pacific (35%) groups than for the total population (15%) between the 1986 and 2013 NZ 506 censuses. This may be explained by increasing housing costs prices, the younger age structure for 507 Māori and Pacific people and lower rates of employment and income levels among these ethnic 508 509 groups (Statistics New Zealand 2016). Other important factors include ethnic differences in intergenerational attitudes to home ownership Statistics New Zealand 2016) and institutionalised 510 racism (Houkamau and Sibley, 2015). Data from the 2002/3 New Zealand Health Survey found that 511 the odds of Māori experiencing racism in the context of housing was 13 times higher than NZ 512 513 Europeans (Harris et al. 2006). Decreasing owner-occupation pushes groups into rental accommodation, insecure by nature and therefore related to residential mobility. A recent survey 514 found that Maori (58%) and Pacific (71%) peoples were more likely to be renters than Asian (41%) or 515 NZ Europeans (27%). To address the issues raised here, future research should assess the impact of 516 517 transitions within and between tenures on ethnic differences in CVD as well as exploring whether and why propensity to migrate varies between ethnic groups. 518

In addition to these confounding factors, it is worth drawing out a final key point of interest from these data. Despite the relative disadvantage of Māori populations who generally have some of the highest probabilities of CVD, the patterning of health for Maori is closely aligned to the experiences of the NZEO. This contrasts with the similarities in the patterning to probabilities of CVD for Pacific, Indian and Other Asian groups. We may speculate that the similarities in the distribution of risk of

524 CVD between these two sets of ethnic groups are related to wider migration and settlement patterns in NZ. Pacific, Indian and Other Asian populations are more likely to comprise recent migrants whose 525 health may follow from their place of origin or are not yet similarly susceptible to the determinants 526 influencing Māori and NZEO health. The similarities between Māori and NZEO groups on the one 527 528 hand, and Pacific, Indian and Other Asian on the other, may therefore be attributed to longevity in NZ and the resulting gradual convergence between cultural and socio-political heritages. As we were 529 unable to exclude (recent) international migrants from the cohort, a common practice in research into 530 selective migration and health (e.g. Norman et al., 2005), this cannot be further tested. However, 531 future work should explore how the influence of residential mobility and deprivation mobility on 532 health may not only vary between ethnic groups in terms of the magnitude of the influence, but also 533 may vary according to length of residence in a country. Such work would build on literatures 534 535 exploring the 'healthy migrant effect' and wider international migration (e.g. Silventoinen et al., 2008; Norredam et al., 2013; Blair and Schneeberg, 2014), rather than internal migration or residential 536 537 mobility.

We have shown that the relationship between residential mobility and risk of CVD varies notably 538 between ethnic groups. However, much of this variation is attributable to the contrasting deprivation 539 540 experiences of different ethnic groups in NZ, evident in the attenuating influence of baseline deprivation circumstances on the odds of CVD by ethnic group, the consistent deprivation gradient in 541 probability of CVD for stayers, and the varying probabilities of CVD for different ethnic groups 542 according to the nature of the move. It is apparent that while residential mobility is an important 543 determinant of CVD in NZ, as was found in the Auckland study, the extent of the influence will vary 544 by ethnic group according to their deprivation experiences. Further differences may also arise if ethnic 545 groups are differentiated by sex as gendered differences in risk of CVD have been determined in the 546 literature (Mieres 2005, Maas and Appelman 2010, Mosca et al., 2011; Brunner, 2016). There may 547 548 also be gendered differences in migration propensities between ethnic groups. Future work should investigate whether gendered differences in risk of CVD interact with possible gendered differences 549 550 in propensity to migrate by ethnic group.

551 Despite the strengths of this study, particularly in the value of the dataset used, there are a number of 552 limitations. Firstly, we are not able to fully disentangle the complexities of the relationship between residential mobility and health in the absence of richer socioeconomic data on the participants 553 included. However, deprivation acts as a good proxy for individual-level socioeconomic data and 554 555 reveals much as to the socially graded risk of CVD and how this varies between ethnic groups. Secondly, we are not able to account for certain factors such as access to healthcare or cultural 556 differences influencing residential mobility patterns. In the case of the latter, it is important to 557 recognise that we are not necessarily comparing like-for-like when looking at different ethnic groups. 558 Relatedly, we must ask whether comparisons between movers and stayers are not necessarily 559 comparing like-for-like: are differences in health outcomes the result of mover or stayer status, or 560 merely an 'artefact of differences in their demographic composition' (Green et al., 2015: 30). While 561 562 the distinct characteristics of mobile groups compared to immobile groups are the basis of theories of health-selective migration, the inherent bias in the data is problematic (note the different composition 563 of movers compared to stayers in Table 1). 564

Green et al. (2015) note that this inherent bias is rarely adequately accounted for in migratory 565 research. To overcome this bias, they advocate the use of 'matching', comparing the change in status 566 567 of one group (e.g. the migration event) with the manually changed status of an alternative control group. Using this pseudo-experimental design, the authors of the study find that migration, regardless 568 of the nature of the move, increased the likelihood that an individual reported poor health. Thus, while 569 the process of matching might help reduce selection bias in the data given the contrasting 570 demographic characteristics of movers compared to stayers, the results of their study are similar to 571 those reported here. Namely, probability of CVD is greater for movers compared to stayers, regardless 572 of the nature of the move. Although this reflects a limitation of the study, our interpretation of the 573 results are still significant. 574

We must look to discussions of health-selective migration to expand on these results. How confident can we be that there is a causal relationship between residential mobility and risk of CVD? The findings presented in this paper contrast with some of the wider literature on migration and health

which finds that migrants, or at least younger migrants, are in better health than their stable 578 579 counterparts (Bentham, 1988; Larson et al., 2004). On the one hand, this may reflect the neglect of 'malign migration' in the literature, something that has also been explored in terms of the 'drift' 580 hypothesis in research exploring mental health and selective migration (see Curtis et al., 2006; De 581 582 Verteuil et al., 2007). The heightened risk of CVD for marginalised minority groups in more deprived areas may be attributed to higher rates of residential mobility. Future research should examine the 583 frequency of moves and the deprivation trajectory of these moves over time to address this issue. On 584 the other hand, the health outcome may be important in assessing the influence of health-selective 585 migration or residential mobility on health inequalities in a population, as is the nature of the move 586 itself in terms of changing deprivation. It is possible that movers may have a heightened susceptibility 587 to certain morbidities such as CVD as a consequence of the move itself. Apart from not having 588 589 experienced a CVD event by the start of the study period, the sequencing of the CVD and migration events are not accounted for here. Thus, for different ethnic groups in NZ, are CVD events the reason 590 for the move (for informal care, for example), are CVD events associated with the move (relating to 591 the stress of moving), or are certain characteristics of movers associated with a higher risk of CVD 592 593 (see forthcoming research)?

Notwithstanding these limitations, this study clearly identifies a number of fruitful avenues for future research. Further, ethnic inequalities in CVD are a major policy concern in NZ, and of international relevance given the existence of these inequalities in countries across the world. The policy implications of this study are clear. Residentially mobile Māori, Pacific and South Asian populations who already have a heightened risk of CVD living in more deprived areas must be the focus of policies aiming to reduce inequalities in CVD within NZ. Moreover, healthcare providers must effectively engage with those mobile vulnerable groups if health inequalities are to reduce.

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Total	Stayers	Movers	Total
	(n =950,151 45.7%)	(n = 1,127,319 54.3%)	(n = 2,077,470)
CVD event			
Yes	75,263 (7.9%)	78,867 (7.0%)	154,130 (7.4%)
No	874,888 (92.1%)	1,048,452 (93.0%)	1,923,340 (92.6%)
Gender			
Male	460,004 (48.4%)	532,608 (47.2%)	992,612 (47.8%)
Female	490,147 (51.6%)	594,711 (52.8%)	1,084,858 (52.2%)
Age			
30-44	333,784 (35.1%)	581,225 (51.6%)	915,009 (44.0%)
45-54	242,051 (25.5%)	251,287 (22.3%)	493,338 (23.7%)
55-64	191,279 (20.1%)	159,863 (14.2%)	351,142 (16.9%)
65-74	119,198 (12.5%)	83,915 (7.4%)	203,113 (9.8%)
75-85	63,839 (6.7%)	51,029 (4.5%)	114,868 (5.5%)
Ethnic			
Māori	65,741 (6.9%)	111,876 (9.9%)	177,617 (8.5%)
Pacific	49,620 (5.2%)	61,641 (5.5%)	111,261 (5.4%)
Indian	22,716 (2.4%)	32,000 (2.6%)	54,716 (6.5%)
Other Asian	61,759 (6.5%)	67,166 (6.0%)	128,961 (6.2%)
NZEO	750,279 (79.0%)	854,636 (75.8%)	1,604,915 (77.3%)
Baseline deprivation			
Q1 – least deprived	235,253 (24.8%)	243,123 (21.6%)	478,376 (23.0%)
Q2	206,990 (21.8%)	235,474 (20.9%)	442,464 (21.3%)
Q3	186,050 (19.6%)	222,702 (19.8%)	408,752 (19.7%)
Q4	169,273 (17.8%)	220,189 (19.5%)	389,462 (18.7%)
Q5 – most deprived	152,585 (16.1%)	205,831 (18.3%)	358,416 (17.3%)
Of movers:			

Table 1. Demographics of movers and stayers aged 30-85 years in New Zealand

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To less deprived area 374,467 (33.2%) Moved within same level 421,114 (37.4%) To more deprived area 331,738 (29.4%) Deprivation transitions 111,072 (9.9%) Moto Q1 111,072 (9.9%) Into Q1 133,457 (11.8%) Out of Q1 118,654 (10.5%) Within Q2-Q4 460,532 (40.9%) Out of Q5 114,158 (10,1%) Into Q5 97,773 (8.7%) Within Q5 91,673 (8.1%) Of stayers: 91,673 (8.1%) Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Deprivation change	
Moved within same level 421,114 (37,4%) To more deprived area 331,738 (29,4%) Deprivation transitions Within Q1 111,072 (9,9%) Into Q1 133,457 (11.8%) Out of Q1 118,654 (10.5%) Within Q2 Q4 460,532 (40.9%) Out of Q5 114,158 (10,1%) Into Q5 97,773 (8,7%) Within Q5 91,673 (8,1%) Of stayers: Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	To less deprived area	374,467 (33.2%)
To more deprived area 331,738 (29.4%) Deprivation transitions 111.072 (9.9%) Mithin Q1 133,457 (11.8%) Out of Q1 133,457 (11.8%) Out of Q1 118,654 (10.5%) Within Q2-Q4 460,532 (40.9%) Out of Q5 114,158 (10,1%) Into Q5 97,773 (8.7%) Within Q5 91,673 (8.1%) Of stayers: 91,673 (8.1%) Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Moved within same level	421,114 (37.4%)
Deprivation transitions 111.072 (9.9%) Within Q1 113.457 (11.8%) Out of Q1 133.457 (11.8%) Out of Q1 118,654 (10.5%) Within Q2-Q4 460.532 (40.9%) Out of Q5 114,158 (10.1%) Into Q5 97,773 (8.7%) Within Q5 91,673 (8.1%) Of stayers: 91,673 (8.1%) Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	To more deprived area	331,738 (29.4%)
Within Q1 111.072 (9.9%) Into Q1 133,457 (11.8%) Out of Q1 118,654 (10.5%) Within Q2-Q4 460,532 (40.9%) Out of Q5 114,158 (10.1%) Into Q5 97,773 (8,7%) Within Q5 91,673 (8.1%) Of stayers: 91,673 (8.1%) Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Deprivation transitions	
Into Q1 133,457 (11.8%) Out of Q1 118,654 (10.5%) Within Q2-Q4 460,532 (40.9%) Out of Q5 114,158 (10.1%) Into Q5 97,773 (8.7%) Within Q5 91,673 (8.1%) Of stayers: 91,673 (8.1%) Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Within Q1	111,072 (9.9%)
Out of Q1 118,654 (10.5%) Within Q2-Q4 460,532 (40.9%) Out of Q5 114,158 (10.1%) Into Q5 97,773 (8.7%) Within Q5 91,673 (8.1%) Of stayers: 91,673 (8.1%) Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Into Q1	133,457 (11.8%)
Within Q2-Q4 460,532 (40.9%) Out of Q5 114,158 (10.1%) Into Q5 97,773 (8.7%) Within Q5 91,673 (8.1%) Of stayers: 91,673 (8.1%) Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Out of Q1	118,654 (10.5%)
Out of Q5 114,158 (10,1%) Into Q5 97,773 (8,7%) Within Q5 91,673 (8,1%) Of stayers: 91,673 (8,1%) Stable Q1 - least deprived 235,253 (24,8%) Stable Q2 206,990 (21,8%) Stable Q3 186,050 (19,6%) Stable Q4 169,273 (17,8%) Stable Q5 - most deprived 152,585 (16,1%)	Within Q2-Q4	460,532 (40.9%)
Into Q5 97,773 (8.7%) Within Q5 91,673 (8.1%) Of stayers: Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Out of Q5	114,158 (10.1%)
Within Q5 91,673 (8.1%) Of stayers: 235,253 (24.8%) Stable Q1 – least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 – most deprived 152,585 (16.1%)	Into Q5	97,773 (8.7%)
Of stayers: Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Within Q5	91,673 (8.1%)
Stable Q1 - least deprived 235,253 (24.8%) Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Of stayers:	
Stable Q2 206,990 (21.8%) Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Stable Q1 – least deprived	235,253 (24.8%)
Stable Q3 186,050 (19.6%) Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Stable Q2	206,990 (21.8%)
Stable Q4 169,273 (17.8%) Stable Q5 - most deprived 152,585 (16.1%)	Stable Q3	186,050 (19.6%)
Stable Q5 – most deprived 152,585 (16.1%)	Stable Q4	169,273 (17.8%)
	Stable Q5 – most deprived	152,585 (16.1%)

Table 2. Population by ethnic group and baseline deprivation quintile

	Q1 Least	Q2	Q3	Q4	Q5 Most
	deprived				deprived
Māori	12,535 (7.1%)	18,181 (10.2%)	26,096 (14.7%)	41,383 (23.3%)	79,422 (44.7%)
Pacific	4,992 (4.5%)	7,889 (7.1%)	12,150 (10.9%)	23,077 (20.7%)	63,153 (56.8%)
Indian	7,341 (13.4%)	9,330 (17.1%)	10,850 (19.8%)	14,777 (27.0%)	12,418 (22.7%)
Other Asian	28,917 (22.4%)	29,455 (22.8%)	26,286 (20.4%)	25,199 (19.5%)	19,104 (14.8%)
NZEO	424,591(26.5%)	377,609 (23.5%)	333,370 (20.8%)	285,026 (17.8%)	184,319 (11.5%)
Total	478,376 (23.0%)	442,464 (21.3%)	408,752 (19.7%)	389,462 (18.7%)	358,416 (17.3%)

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Model description	Model 1 Odds Ratio (95% CI)	Model 2 Odds Ratio (95% CI)	Model 3 Odds Ratio (95% CI)	Model 4 Odds Ratio (95% CI)	Model 5 Odds Ratio (95% CI)
Adjusts for gender, age ethnicity plus:	Mover status	Mover status, baseline deprivation	Deprivation mobility status	Detailed deprivation transitions	Deprivation quintile for stayers
Gender			•		
Female	REF	REF	REF	REF	REF
Male	$1.66^{\circ} (1.64 - 1.68)$	1.66*(1.64 - 1.68)	1.66*(1.64 - 1.68)	$1.66^{*}(1.64 - 1.68)$	$1.66^{*}(1.64 - 1.68)$
Age group					
30-44	0.12*(0.12-0.12)	0.12*(0.12-0.12)	$0.12^{*}(0.12^{-0.12})$	$0.12^{*}(0.12 - 0.12)$	0.12*(0.12-0.12)
45-54	0.42* (0.42 -0.43)	0.42* (0.42 -0.43)	0.42* (0.42 -0.43)	0.42* (0.42 -0.43)	0.43*(0.42 -0.43)
55-64	REF	REF	REF	REF	REF
65-74	$2.41^{*}(2.37 - 2.44)$	2.38*(2.34 - 2.42)	2.40*(2.37 - 2.44)	2.39*(2.35-2.43)	2.39*(2.36-2.43)
75-85	$5.54^{*}(5.45 - 5.63)$	5.43*(5.34-5.52)	$5.54^{*}(5.44 - 5.63)$	$5.48^{*}(5.39 - 5.57)$	5.48*(5.39-5.58)
Ethnicity			1		
NZEO	REF	REF	REF	REF	REF
Māori	2.25*(2.21-2.30)	$1.97^{*}(1.93-2.01)$	2.26*(2.21-2.30)	2.05*(2.01-2.09)	2.15*(2.10-2.19)
Pacific	$1.63^{*}(1.59 - 1.67)$	1.38^{*} $(1.35 - 1.42)$	$1.64^{*}(1.60 - 1.68)$	$1.47^{*}(1.43 - 1.51)$	1.53^{*} $(1.49 - 1.57)$
Indian	1.21*(1.17 - 1.26)	$1.14^{*}(1.10 - 1.19)$	7 1.21* (1.17 - 1.26)	1.17*(1.12 - 1.22)	1.19*(1.15 - 1.24)
Other Asian	0.56*(0.54 - 0.58)	0.55* (0.54 - 0.57)	0.56*(0.54 - 0.58)	$0.56^{*} (0.54 - 0.57)$	$0.56^{*}(0.54 - 0.58)$
Mover status					
Stayer	REF	REF			
Mover	$1.26^{*}(1.25 - 1.28)$	1.26*(1.24-1.27)			
Baseline deprivation (NZDe	:p2006)				
Q1(least deprived)		REF			
Q2		$1.14^{*}(1.12 - 1.16)$			
Q3		1.26*(1.24-1.29)			
Q4		1.39*(1.37-1.42)			
Q5		1.58*(1.55-1.61)			
Deprivation mobility status					
Stayer			REF		
Moves up			$1.29^{*}(1.27 - 1.31)$		
Moves w/in		~	1.23*(1.21-1.25)		
Moves down			$1.28^{*}(1.26 - 1.30)$		
Deprivation transitions (det:	ailed moves between quin	tiles)			
Stayer				REF	
Within Q1				$(19.0 - 68.0) \times 10.00$	

774 Table 3. Binary logistic regression modelling CVD events in NZ adult population

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REF 0.65* (0.64 – 0.67) 0.73* (0.72 – 0.75) 0.81* (0.79 – 0.82) 0.89* (0.87 – 0.90) 0.94* (0.92 – 0.96)	
$\begin{array}{c} 1.08^{*} \left(1.05 - 1.11 \right) \\ 1.06^{*} \left(1.03 - 1.08 \right) \\ 1.26^{*} \left(1.24 - 1.28 \right) \\ 1.55^{*} \left(1.51 - 1.58 \right) \\ 1.57^{*} \left(1.48 - 1.56 \right) \\ 1.71^{*} \left(1.66 - 1.76 \right) \end{array}$	
S.	21 mg
	COLOR
0 Q1 0f Q1 02 4 02 4 0 Q5 in Q5 in Q5 le Q1 le Q2 le Q3 le Q4 le Q5	gnificant ORs are starred: p
Int Out o Within Out o Int Withi Stable deprivation Stabl Stabl Stabl Stabl	775 Note: statistically si

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776 Figure captions

- 777 Figure 1. Population eligibility flow chart
- Figure 1 Probability of CVD (%) by mover status, stratified by ethnic group (adjusting for age,
 gender, [and ethnicity])
- Figure 2 Probability of a patient having CVD (%) by baseline deprivation, stratified by ethnic group
 (adjusting for mover status, age, gender, [and ethnicity])
- Figure 3 Probability of CVD (%) by deprivation mobility status, stratified by ethnic group (adjusting
 for age gender, [ethnicity])
- Figure 4 Probability of CVD (%) by detailed deprivation transition, stratified by ethnic group
 (adjusting for age, gender, [ethnicity])
- Figure 5 Probability of CVD (%) by stable deprivation for stayers compared to movers, stratified by
 ethnic group (adjusting for age, gender, [ethnicity])
- 788
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- 792

793 Figure 1



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Highlights

- Evidence suggests residential mobility contributes to widening health inequalities
- Little is known about ethnic variations in residential mobility
- We investigate associations between residential mobility and CVD in New Zealand
- Much of the ethnic variation is explained by contrasting deprivation experiences
- We also show a deprivation gradient in CVD risk among stayers as well as movers

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